Multi-sensor Information Fusion Steady-State Kalman Estimator for Systems with System Errors and Sensor Errors

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Abstract

In this paper, a multi-sensor information fusion steady-state Kalman estimator for discrete time stochastic linear systems with system errors and sensor errors is presented. Gevers-Wouters(G-W) algorithm is used in this paper. Steady-state Kalman estimator is presented in this paper avoids the complex Diophantine equation, and it is based on the ARMA model to compute the steady-state Kalman estimators gain, further the Lyapunov equation is used to estimate the variance matrix and covariance matrix of estimation error. So this algorithm can obviously reduce the computational burden. In order to improve the estimation accuracy, the multi-sensor information fusion method is adopted. The fusion method includes weighted measurement fusion, weighted by scalars and the covariance intersection fusion. Under the linear minimum variance optimal information fusion criterion, the calculation formula of optimal weighting coefficients have been given in order to realize scalars weighted. To avoid the calculation of cross-covariance matrices, another distributed fusion filter is also presented by using the covariance intersection fusion algorithm, which can reduce the computational burden. And the relationship between the accuracy and the computation complexities among the three fusion algorithm are analyzed. A simulation example of the target tracking controllable system with two sensors shows its effectiveness and correctness.

Keywords: multi-sensor information fusion; the linear minimum variance; the steady-state estimator; system errors and sensor errors; Gevers-Wouters algorithm

1. Introduction

Classic Kalman filtering is a time varying recursive filter. The optimal Kalman filter is required to compute the gain matrix, the variance matrix and covariance matrix of estimation error at every moment based on Riccati equation, which brings a large computational burden. The steady state Kalman estimators based on modern time series analysis method, which is based on the ARMA model to obtain the steady-state Kalman estimators gain. And the Lyapunov equation is used to estimate the variance matrix and the covariance matrix of the estimation error. This method completely avoids the Riccati equation, and uses the the ARMA model to replace the Riccati equation. The method can obviously reduce the computational burden, and the method can be used to design the self-tuning steady-state Kalman estimators of unknown noise statistics system.

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In order to improve the estimation accuracy of the sensor, multi-sensor information fusion is presented. Multi-sensor information fusion is how to combine the local state estimators or local measurements which is obtained from each sensor, to obtain the fused state estimator. And its accuracy is higher than that of each local estimator. There are two common types of information fusion method. One is state fusion, the other is measurement fusion. The state fusion method is also divided into centralized Kalman filtering and distributed Kalman filtering. Although the centralized Kalman filter can obtain the global optimal fusion state estimation in theory, and it has the disadvantage of large computation burden and poor fault tolerance, and the distributed Kalman filtering information fusion can overcome these drawbacks. Distributed fusion Kalman filtering under linear minimum variance rules has three information fusion algorithms: the matrix weighted, scalar weighted and diagonal matrix weighted. The distributed fusion estimation needs to calculate the cross-covariances of local estimate. However, in many theoretical and application problems, the cross-covariances are unknown, or it is very difficult to compute the cross-covariances, or can’t calculate the cross-covariances. If the cross-covariances are neglected, it will lead the increase of the variance of the local estimator error, even the divergence of the estimator.

In order to overcome the disadvantages and limitations, Jeffrey K. Uhlman proposed the covariance intersection information fusion method. It is also a distributed information fusion method, and it completely avoids the unknown cross-covariance identification and calculation. It can deal with the fusion estimation problem for the system with unknown covariance. This algorithm avoids computing the covariance, so it can reduce the computational burden. Further the algorithm can be used to fusion estimator for nonlinear systems.

In this paper, a multi-sensor information fusion steady-state Kalman estimator for discrete time stochastic linear systems with system errors and sensor errors is presented. Firstly, steady-state Kalman estimator is presented in this paper can obviously reduce the computational burden. Secondly, the fusion method includes weighted measurement fusion, weighted by scalars and the covariance intersection fusion. The accuracy of above three kinds of weighted fusion steady-state Kalman estimator from high to low is weighted measurement fusion, scalars weighted and the covariance intersection fusion. But the computational burden is on the contrary, the weighted measurement fusion estimator has a large computational burden. And covariance intersection fusion avoids solving cross-covariance matrices. So it has the minimal computational burden. Lastly, A simulation example of the target tracking controllable system with two sensors shows its effectiveness and correctness.

The main structure of this paper is as follows: Problem formulation is given in Section 2. The Local optimal steady-state estimator is obtained in Section 3. In Section 4 multi-sensor information fusion optimal steady-state estimator is presented. A simulation example with 2-sensor is given is Section 5. In Section 6 the conclusions of this paper are given.

2. Problem Formulation

Consider the multi-sensor discrete time time-invariant stochastic linear system with the same observation array, the system errors and sensor errors

\[
x(t+1) = \Phi x(t) + F w(t) + U d(t)
\]

\[
d(t+1) = d(t) + \zeta(t)
\]

\[
y_i(t) = H x(t) + v_i(t) + e_i(t)
\]

(3)
\( e_i(t + 1) = e_i(t) + \xi_i(t), \quad i = 1, \ldots, L \)

where \( x(t) \in \mathbb{R}^n \) is the state of the system, \( y_i(t) \in \mathbb{R}^m \) is the measurement of the \( i \)th sensor subsystem, \( \Phi, \Gamma, U, H \) is the suitable dimensional matrix respectively, \( d(t) \in \mathbb{R}^r \) is system errors and \( e_i(t) \in \mathbb{R}^m \) is sensor errors. \( \xi_i(t), \quad i = 1, 2, \ldots, L \) are independence white noises with zero mean, and they and \( w(t) \in \mathbb{R}^r, \quad v_i(t) \in \mathbb{R}^m, \quad i = 1, \ldots, L \) are independence. And the subscript \( i \) is the \( i \)th sensor of all the sensor, \( L \) is the number of sensor in the system.

**Assumption 1** \( w(t) \in \mathbb{R}^r \) and \( v_j(t) \in \mathbb{R}^m, \quad j = 1, \ldots, K \) independence white noises with zero mean and covariance are \( Q_w \) and \( Q_v \) individually

\[
E \left[ \begin{bmatrix} w(t) \\ v_i(t) \end{bmatrix} \right] = \begin{bmatrix} Q_w \\ 0 \end{bmatrix}, \quad \delta_{wk} = 0(t \neq k)
\]

where the superscript \( T \) is the transpose, \( E \) is the expectation, \( \delta_{wk} = 1, \quad \delta_{k} = 0(t \neq k) \).

**Assumption 2** \( (\Phi, H) \) is completely observable pair, and \( (\Phi, \Gamma) \) is completely stable pair.

**Assumption 3** The initial time \( t_0 = -\infty \).

Steady-state estimation problem is based on the measurement \( (y_i(t + N), y_i(t + N - 1), \ldots) \), to obtain the linear minimum variance state \( \hat{x}_i(t | t + N), \quad i = 1, 2, \ldots, L \). For \( N = 0, \quad N > 0 \) or \( N < 0 \), we named it as steady-state filtering, smoothing or predictor. Further weighted measurement fusion and distributed optimal information fusion steady-state estimation \( \hat{x}_0(t | t + N) \) is obtained, it consists of weighted local steady-state estimators.

### 3. Local Optimal Steady-State Estimator

For the system (1)–(4), we can get the following augmentation system

\[
\bar{x}_i(t + 1) = \bar{\Phi}_i \bar{x}_i(t) + \bar{\Gamma}_i \bar{w}_i(t)
\]

(6)

\[
y_i(t) = \bar{H} \bar{x}_i(t) + v_i(t), \quad i = 1, \ldots, L
\]

(7)

where

\[
\bar{x}_i(t) = \begin{bmatrix} x(t) \\ d(t) \\ e_i(t) \end{bmatrix}, \quad \bar{w}_i(t) = \begin{bmatrix} w(t) \\ \xi_i(t) \end{bmatrix}, \quad \bar{\Phi}_i = \begin{bmatrix} \Phi & U & 0 \\ 0 & I_w & 0 \\ 0 & 0 & \beta_i \end{bmatrix}, \quad \bar{\Gamma}_i = \begin{bmatrix} \Gamma & 0 & 0 \\ 0 & I_w & 0 \\ 0 & 0 & I_w \end{bmatrix}, \quad \bar{H} = \begin{bmatrix} H & 0 & I_w \end{bmatrix}
\]

The state \( x(t) \) is the firstly \( n \)th components for the system (6) and (7), so, the original problem converts to the fusion steady-state estimator problem for the system (7) and (8).

The system noise \( \bar{w}_i(t) \) and observation noise \( v_i(t) \) are correlated, and

\[
E \left[ \begin{bmatrix} \bar{w}_i(t) \\ v_i(t) \end{bmatrix} \right] = \begin{bmatrix} Q_{\bar{w}_i(t)} \\ 0 \end{bmatrix}
\]

(9)

where \( Q_{\bar{w}_i(t)} = Q_v_i = \text{diag}(Q_v_i, Q_v_i, \ldots), \quad Q_{v_i(t)} = \text{diag}(Q_v_i, Q_v_i, 0), \quad i \neq j \).

From (6) and (7) having

\[
y_i(t) = \bar{H} (I_w - q^{-1} \bar{\Phi}_i) \bar{\Gamma}_i q^{-1} \bar{w}_i(t) + v_i(t)
\]

(10)

where \( q^{-1} \) is the lag operator for the unit, introducing left decomposition
\[ \dot{H}(I_n - q^{-1}\hat{\Phi})^{-1}\dot{T}q^{-1} = A^{(i+1)}(q^{-1})B^{(i)}(q^{-1}) \]

(11)

where \( A^{(i)}(q^{-1}) \) and \( B^{(i)}(q^{-1}) \) is matrix polynomial, it is shaped as \( X^{(i)}(q^{-1}) = x_i^{(0)} + x_i^{(0)}q^{-1} + \cdots + x_n^{(0)}q^{-n} \), and \( A_0^{(i)} = I_n, B_0^{(i)} = 0 \), letting \( X^{(i+1)}(q^{-1}) = (X^{(i)}(q^{-1}))^{-1} \).

Substituting (11) into (10) yields

\[ A^{(i+1)}(q^{-1})y(t) = B^{(i)}(q^{-1})\bar{v}(t) + A^{(i+1)}(q^{-1})v_i(t) \]

(12)

For \( (A^{(i)}(q^{-1}), B^{(i)}(q^{-1})) \) is left and Assumption 1, yielding the ARMA innovation model

\[ A^{(i+1)}(q^{-1})y(t) = D^{(i)}(q^{-1})\varepsilon(t) \]

(13)

where innovation \( \varepsilon(t) \in \mathbb{R}^{n_o} \) is white noises with zero mean, and its covariance is \( Q_{\varepsilon} \).

\[ D^{(i)}(q^{-1}) \] is stable, \( D_{0}^{(i)} = I_n \), and having

\[ D^{(i)}(q^{-1})\varepsilon(t) = B^{(i)}(q^{-1})\bar{v}(t) + A^{(i+1)}(q^{-1})v_i(t) \]

(14)

\( D^{(i)}(q^{-1}) \) and \( Q_{\varepsilon} \) are computed through G-W algorithm.

**Lemma 1** For the system (6) and (7) under the Assumption 1-4, the \( i \)th sensor subsystem has the local optimal steady-state Kalman filtering equations:

\[ \dot{x}_i(t+1|t+1) = \Psi_{\beta_i}\dot{x}_i(t|t) + [I_n - K_{\beta_i}\dot{H}]y_i(t) + K_{\beta_i}y_i(t+1) \]

(15)

\[ \Psi_{\beta_i} = [I_n - K_{\beta_i}\dot{H}]\bar{\Phi}_i \]

(16)

where \( \Psi_{\beta_i} \) is a stable matrix, the filtering gain \( K_{\beta_i} \) is as follows:

\[ K_{\beta_i} = \begin{bmatrix} \dot{H} & \hat{\Phi} & \Phi \Phi^{-1} & I_{m_n} - Q_{\varepsilon} \\ M^{(i)}_{\beta_i} & \cdots & \cdots \\ \ddots & \cdots & \cdots \\ \hat{\Phi}\Phi^{-1} & \cdots & \cdots & M^{(i)}_{\beta_{i-1}} \end{bmatrix} \]

(17)

The pseudo inverse \( X^* \) of the matrix \( X \) is defined as \( X^* = (X^TX)^{-1}X^T \), \( \beta_i \) is index of correlation of \( \hat{\Phi}_i, \dot{H} \), the coefficient matrix \( M^{(i)}_{\beta_i} \) can be recursively computed as

\[ M^{(i)}_{\beta_i} = -A_{\beta_i}M^{(i)}_{\beta_{i-1}} - \cdots - A_{n_{\beta_i}}M^{(i)}_{k=n_{\beta_i}} + D_{\beta_i} \]

(18)

where \( M^{(i)}_{\beta_i} = I_{n_i}, M^{(i)}_{k} = 0(k < 0), D_{\beta_i} = 0(k > n_{\beta_i}) \).

Steady local filtering error covariance matrix \( P_i \) satisfies the Lyapunov equation

\[ P_i = \Psi_{\beta_i}P_i\Psi_{\beta_i}^T + [I_n - K_{\beta_i}\dot{H}]\bar{\Phi}Q_{\varepsilon}\bar{\Phi}^T[I_n - K_{\beta_i}\dot{H}]^T + K_{\beta_i}Q_{\varepsilon}K_{\beta_i}^T \]

(19)

the cross covariance \( P_i^0 \) between any two local filtering satisfies Lyapunov equation:

\[ P_i^0 = \Psi_{\beta_i}P_i\Psi_{\beta_i}^T + [I_n - K_{\beta_i}\dot{H}][\bar{\Phi}Q_{\varepsilon}\bar{\Phi}^T][I_n - K_{\beta_i}\dot{H}]^T + K_{\beta_i}Q_{\varepsilon}K_{\beta_i}^T \]

(20)

**Lemma 2** For the system (6) and (7) under the Assumption 1-4, the \( i \)th sensor subsystem has the local optimal steady-state Kalman prediction equations:

\[ \dot{x}_i(t+1|t) = \Psi_{\beta_i}\dot{x}_i(t|t) + K_{\beta_i}y_i(t) \]

(21)

\[ \Psi_{\beta_i} = \bar{\Phi}_i - K_{\beta_i}\dot{H} \]

(22)
where \( \Psi \) is a stable matrix, the prediction gain \( K \) is as follows:

\[
K = \begin{bmatrix}
\tilde{H} \\
\tilde{H}\Phi \\
\vdots \\
\tilde{H}\Phi^{k-1}
\end{bmatrix} \begin{bmatrix}
M^{(1)}_i \\
M^{(2)}_i \\
\vdots \\
M^{(k)}_i
\end{bmatrix}
\]

(23)

where the coefficient matrix \( M^{(k)}_i \) can be recursively computed by (18).

Steady local prediction error covariance matrix \( \Sigma_i \) satisfies the Lyapunov equation

\[
\Sigma_i = \Psi_i \Sigma_i \Psi_i^T + \tilde{F}Q_{ni} \tilde{F}^T + K_{pi}Q_{pi} K_{pi}^T
\]

(24)

the cross covariance \( \Sigma_{ij} \) between any two local prediction satisfies Lyapunov equation:

\[
\Sigma_{ij} = \Psi_i \Sigma_{ij} \Psi_j^T \quad i \neq j
\]

(25)

**Lemma 3** For the system (6) and (7) under the Assumption 1-4, the \( i \)th sensor subsystem has the local \((-N)\)-step-ahead optimal steady-state Kalman prediction equations:

\[
\hat{x}_i(t | t + N) = \tilde{H}^{-N-1} \hat{x}_i(t + N + 1 | t + N) \quad N \leq 2, \quad i = 1, \ldots, L
\]

(26)

Steady local filter error covariance matrix \( P_i(N) \) can be calculated by

\[
P_i(N) = \tilde{H}^{-N-1} \Sigma_i \tilde{H}^{-(N+1)} + \sum_{j=0}^{N-2} \tilde{H}^j \Gamma Q \Gamma^T \tilde{H}^j \quad N \leq 2
\]

(27)

**Lemma 4** For system (6) and (7) under the assumption 1-4, the \( i \)th sensor subsystem has the local optimal steady-state Kalman smoothing equations:

\[
\hat{x}_i(t | t + N) = \hat{x}_i(t | t - 1) + \sum_{k=0}^{N} K_{mi} (k) e_i(t + k) \quad i = 1, \ldots, L
\]

(28)

where \( \hat{x}_i(t | t - 1) \) and \( e_i(t) = y_i(t) - \tilde{H} \hat{x}_i(t | t - 1) \) can be computed by Lemma 2.

The smoothing gain \( K_{mi} \) is as follows:

\[
K_{mi}(k) = \Sigma_i \Psi_i^{k+1} \tilde{H}^i Q_{mi}^{-1}
\]

(29)

\[
Q_{mi} = \tilde{H} \Sigma_i \tilde{H}^T + Q_{mi}
\]

(30)

\[
K_{pi} = \tilde{H} \Sigma_i \tilde{H}^T Q_{mi}^{-1}
\]

(31)

\( \Sigma_i \) can be computed by Lemma 2.

Steady local smoothing error covariance matrix \( P_i(N) \) satisfies the Lyapunov equation

\[
P_i(N) = \Sigma_i - \sum_{k=0}^{N} K_{mi}(k) Q_{mi} K_{mi}^T(k)
\]

(32)

the cross covariance \( P_{ij}(N) \) between any two local filtering satisfies Lyapunov equation:
\[ P_y(N) = \Sigma_y = -\sum_{r=0}^{N} K_m(r) \bar{A} \Sigma_y \bar{A}^\top - \sum_{r=0}^{N} \sum_{s=0}^{r} \Sigma_y \bar{A} K_m(s) \bar{A} \Sigma_y + \sum_{r=0}^{N} \sum_{s=0}^{r} K_m(r) E_y(r,s) K_m(s)^\top \]

where \( \Sigma_y \) can be calculated by (25).

\[ E_y(r,s) = \bar{A} \Sigma_y \bar{A}^\top + \sum_{t=0}^{\min(r,s)} \bar{A} \Sigma_y \bar{A}^{t-1} \begin{bmatrix} Q_{r,s} & 0 \\ 0 & -K_{r,s} \end{bmatrix} \]

\[ \begin{bmatrix} \bar{P} \\ \bar{Q}_{r,s} \end{bmatrix} = \begin{bmatrix} \bar{P} \\ -K_{r,s} \end{bmatrix} \]

\[ \Sigma_y \]

\[ \bar{P} = \bar{H} \Sigma_y \bar{H}^\top \]

\( i \neq j, \min(r,s) > 0 \)

If \( \min(r,s) = 0 \), then having

\[ E_y(0,0) = \bar{H} \Sigma_y \bar{H}^\top \]

\[ E_y(r,0) = \bar{H} \Sigma_y \bar{H}^\top \]

\[ E_y(0,s) = \bar{H} \Sigma_y \bar{H}^\top \]

4. Multi-Sensor Information Fusion Optimal Steady-State Estimator

The system (7) can be written as

\[ \bar{y}^{(o)}(t) = \bar{H}^{(o)} x(t) + \bar{v}^{(o)}(t) \]

\[ \bar{y}^{(o)}(t) = \left[ \sum_{i=1}^{I_m} Q_{+,1}^{-1} \right] \sum_{i=1}^{I_m} Q_{+,1}^{-1} y_i(t) \]

\[ \bar{v}^{(o)}(t) = \left[ \sum_{i=1}^{I_m} Q_{+,1}^{-1} \right] \sum_{i=1}^{I_m} Q_{+,1}^{-1} v_i(t) \]

\[ Q_{v}^{(o)} = \left[ \sum_{i=1}^{I_m} Q_{+,1}^{-1} \right]^{-1} \]

**Theorem 1** For the system (6) and (38) under the Assumption 1-4, global optimality weighted measurement fusion steady-state Kalman filtering equations:

\[ \dot{\bar{x}}^{(o)}(t+1) = \bar{A} \bar{x}^{(o)}(t) + K_f \bar{y}^{(o)}(t) \]

\[ \bar{y}^{(o)}(t) = [I_n - K_f \bar{H}^{(o)}] \bar{\phi} \]

Where \( \bar{\phi} \) is a stable matrix, the filtering gain \( K_f \) is as follows:

\[ K_f = \begin{bmatrix} \bar{H}^{(o)} & \bar{H}^{(o)} \bar{\phi} \\ \vdots & \vdots \\ \bar{H}^{(o)} \bar{\phi}^{\beta-1} & M_{1} \end{bmatrix} \begin{bmatrix} I_m - Q_{v}^{(o)} \bar{\phi} \bar{H}^{(o)} \\ M_{1} \\ \vdots \\ M_{\beta-1} \end{bmatrix} \]

The pseudo inverse \( X^+ \) of the matrix \( X \) is defined as \( X^+ = (X^\top X)^{-1} X^\top \). \( \beta \) is index of correlation of \( (\bar{\phi}, \bar{H}^{(o)}) \), the coefficient matrix \( M_k \) can be recursively computed as

\[ M_k = -A_k M_{k-1} - \cdots - A_1 M_{k-n_k} + D_k \]

where \( M_0 = I_n, M_k = 0(k < 0), D_k = 0(k > n_k) \).

Optimality weighted measurement fusion filtering error covariance matrix \( P_0 \) satisfies the Lyapunov equation.
\[ P_0 = \Psi_j P_0 \Psi_j^T + [I_n - K_j \bar{H}^{(o)}_j] \bar{F} Q_{\pi} \bar{F}^T [I_n - K_j \bar{H}^{(o)}_j]^T + K_j Q^{(o)}_j K_j^T \]  

(46)

And global optimality weighted measurement fusion steady-state Kalman prediction equations:

\[ \dot{x}^{(o)}(t+1|t) = \Psi_p \dot{x}^{(o)}(t|t-1) + K_p y(t) \]  

(47)

\[ \Psi_p = \bar{\Phi} - K_p \bar{H}^{(o)} \]  

(48)

where \( \Psi_p \) is a stable matrix, the prediction gain \( K_p \) is as follows:

\[
K_p = \begin{bmatrix}
\bar{H}^{(o)}_1 & \bar{H}^{(o)}_2 & \cdots & \bar{H}^{(o)}_{p-1}
\end{bmatrix}
\begin{bmatrix}
M_1 \\
M_2 \\
\vdots \\
M_{p-1}
\end{bmatrix}
\]  

(49)

where the coefficient matrix \( M_i \) can be recursively computed by (45).

Optimality weighted measurement fusion prediction error covariance matrix \( \Sigma_0 \) satisfies the Lyapunov equation

\[ \Sigma_0 = \Psi_p \Sigma_p \Psi_p^T + \bar{F} Q_{\pi} \bar{F}^T + K_p Q^{(o)}_p K_p^T \]  

(50)

And global optimality weighted measurement fusion steady-state Kalman smoothing equations:

\[ \dot{x}^{(o)}(t|N) = \dot{x}^{(o)}(t|t-1) + \sum_{k=0}^{N} K_m(k)e(t+k) \]  

(51)

\[ K_m(k) = \Sigma_p \Psi_p \bar{H}^{(o)}_k Q_{\pi}^{-1} \]  

(52)

\[ Q_m = \bar{H}^{(o)}_k \Sigma_p \bar{H}^{(o)}_k^T + Q^{(o)}_p \]  

(53)

\[ \Psi_p = \bar{\Phi} - K_p \bar{H}^{(o)} \]  

(54)

where \( K_p \) can be calculated by (49).

Optimality weighted measurement fusion prediction error covariance matrix \( P_s(N) \) satisfies the Lyapunov equation

\[ P_s(N) = \Sigma_0 - \sum_{k=0}^{N} K_m(k) Q_m K_m^T (k) \]  

(55)

**Theorem 2** For the system (6) and (7), under the Assumption 1 -4, the optimal fused steady-state Kalman filtering \( \hat{x}_0(t|t) \) weighted by scalars is given as

\[ \hat{x}_0(t|t) = \sum_{i=1}^{L} \alpha_i \hat{x}_i(t|t) \]  

(56)

Under the linear minimum variance optimal information fusion criterion which minimize the performance index, the optimal weighting coefficients \( \alpha_i, i = 1, 2, \ldots, L \) are given by
\[ [\alpha_1, \ldots, \alpha_L] = \frac{e^T P_u^{-1}}{e^T P_u^{-1} e} \]

(57)

where we define the \( L \times L \) matrix \( P_u = (\text{tr} P_y)_{L \times L}, i, j = 1, 2, \ldots, L \), and \( P_y \) can be calculated by (20), and \( L \times 1 \) row vector \( e = [1 \quad \cdots \quad 1]^T \).

The optimal fused variance matrix is given as

\[ P_\alpha(N) = \sum_{i,j=1}^{L} \alpha_i \alpha_j P_y(N) \]

(58)

and

\[ \text{tr} P_\alpha \leq \text{tr} P_y, \quad i = 1, 2, \ldots, L \]

(59)

**Theorem 3** For the system (6) and (7), under the same conditions, when the variance of \( P_1 \) and \( P_2 \) are known, but the cross covariance \( P_{12} \) is unknown, using the covariance intersection (CI) fusion method, this paper proposes a suboptimal fusion Kalman estimators is as follows:

\[ \hat{x}_{CI}(t | t + N) = P_C \omega P_1^{-1} \hat{x}_1(t | t + N) + (1 - \omega) P_2^{-1} \hat{x}_2(t | t + N) \]

(60)

where \( \omega \in [0, 1] \) and minimizes the performance index

\[ J = \min \omega \text{tr} P_{CI} \]

(61)

and \( P_{CI} \) is defined as

\[ P_{CI} = [\omega P_1^{-1} + (1 - \omega) P_2^{-1}]^{-1} \]

(62)

For the non-linear optimization problems (62), the optimal weights \( \omega^* \) can be obtained by 0.618 method or the Fabonacci method.

5. Simulation Example

Consider 2-sensor discrete-time linear time-invariant stochastic tracking system (6) and (7), where \( T = 0.7 \) is the sampled period, \( A = \begin{bmatrix} 1 & T \\ 0 & 1 \end{bmatrix} \), \( F = \begin{bmatrix} T^2/2 & T \\ T & 1 \end{bmatrix} \), \( H = [1 \ 0] \). And \( w(t) \) is white noises with zero mean and its covariance is \( \sigma_w^2 = 0.1 \). \( v(t) \) and \( w(t) \) are independence noise, \( \sigma_v^2 = 0.1, \sigma_w^2 = 0.5 \).

Steady-state estimation problem is based on the measurement \( y_i(t+N), y_i(t+N-1), \ldots \), to obtain the linear minimum variance state \( \hat{x}_i(t | t + N) \) and optimal information fusion steady-state estimation \( \hat{x}_o(t | t + N) \).
Fig. 1 The position and fusion steady-state weighted measurement fusion filter.

Fig. 2 The velocity and fusion steady-state weighted measurement fusion filter.

Fig. 3 The position and fusion steady-state fusion filter weighted by scalar.

Fig. 4 The velocity and fusion steady-state fusion filter weighted by scalar.

Fig. 5 The position and fusion steady-state the covariance intersection fusion filter.

Fig. 6 The velocity and fusion steady-state the covariance intersection fusion filter.

Fig. 7 The curves of the sum of absolute error curve for local and fusion filters of the position.

Fig. 8 The curves of the sum of absolute error curve for local and fusion filters of the velocity.

- **black**: weighted measurement fusion
- **green**: weighted by the covariance intersection fusion
- **red**: subsystem 1
- **blue**: subsystem 2
The simulation results are shown in Figure 1-Figure 8. Figure 1 to 6 gives the fusion filter weighted measurement fusion, weighted by scalars and the covariance intersection fusion. Figure 7 and Figure 8 are the absolute error curve for the steady-state filter and fusion steady-state filter of position and velocity weighted measurement fusion, weighted by scalars and the covariance. In the figure, the accuracy of the fusion state filter is higher than any of the single sensor. The accuracy of above three kinds of weighted fusion estimator from high to low is the weighted measurement fusion, weighted by scalars, and covariance intersection fusion. But the computational burden is on the contrary, the weighted measurement fusion estimator has a large computational burden. And covariance intersection fusion avoids solving cross-covariance matrices and has the minimal computational burden.

6. Conclusions

In this paper, a multi-sensor information fusion steady-state Kalman estimator for discrete time stochastic linear systems with system errors and sensor errors is presented. Information fusion rule, which adopted in this paper, includes weighted measurement fusion, weighted by scalars and the covariance intersection fusion. The estimation accuracy for the system is greatly improved compared with the single local sensor. Weighted measurement fusion, scalars weighted and the covariance intersection fusion, the accuracy of above three kinds of weighted fusion filtering from high to low. But the computational burden is on the contrary. Fusion filtering weighted by weighted measurement fusion has a large computational burden, and the covariance intersection fusion with minimal computational burden, and it is suitable for real-time applications. The simulation example shows its validity. The algorithm presented in this paper has many advantages.

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